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Effect of Antibiotic Prescription Audit and Feedback on Antibiotic Prescribing in Primary Care A Randomized Clinical Trial

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IMPORTANCE Antibiotics are commonly prescribed in primary care, increasing the risk of antimicrobial resistance in the population.

OBJECTIVE To investigate the effect of quarterly audit and feedback on antibiotic prescribing among primary care physicians in Switzerland with medium to high antibiotic prescription rates.

DESIGN, SETTING, AND PARTICIPANTS This pragmatic randomized clinical trial was conducted from January 1, 2018, to December 31, 2019, among 3426 registered primary care physicians and pediatricians in single or small practices in Switzerland who were among the top 75% prescribers of antibiotics. Intention-to-treat analysis was performed using analysis of covariance models and conducted from September 1, 2021, to January 31, 2022.

INTERVENTIONS Primary care physicians were randomized in a 1:1 fashion to undergo quarterly antibiotic prescribing audit and feedback with peer benchmarking vs no intervention for 2 years, with 2017 used as the baseline year. Anonymized patient-level claims data from 3 health insurers serving roughly 50% of insurees in Switzerland were used for audit and feedback. The intervention group also received evidence-based guidelines for respiratory tract and urinary tract infection management and community antibiotic resistance information. Physicians in the intervention group were blinded regarding the nature of the trial, and physicians in the control group were not informed of the trial.

MAIN OUTCOMES AND MEASURES The claims data used for audit and feedback were analyzed to assess outcomes. Primary outcome was the antibiotic prescribing rate per 100 consultations during the second year of the intervention. Secondary end points included overall antibiotic use in the first year and over 2 years, use of quinolones and oral cephalosporins, all-cause hospitalizations, and antibiotic use in 3 age groups.

RESULTS A total of 3426 physicians were randomized to the intervention (n = 1713) and control groups (n = 1713) serving 629 825 and 622 344 patients, respectively, with a total of 4 790 525 consultations in the baseline year of 2017. In the entire cohort, a 4.2% (95% CI, 3.9%-4.6%) relative increase in the antibiotic prescribing rate was noted during the second year of the intervention compared with 2017. In the intervention group, the median annual antibiotic prescribing rate per 100 consultations was 8.2 (IQR, 6.1-11.4) in the second year of the intervention and was 8.4 (IQR, 6.0-11.8) in the control group. Relative to the overall increase, a -0.1% (95% CI, -1.2% to 1.0%) lower antibiotic prescribing rate per 100 consultations was found in the intervention group compared with the control group. No relevant reductions in specific antibiotic prescribing rates were noted between groups except for quinolones in the second year of the intervention (-0.9% [95% CI, -1.5% to -0.4%]).

CONCLUSIONS AND RELEVANCE This randomized clinical trial found that quarterly personalized antibiotic prescribing audit and feedback with peer benchmarking did not reduce antibiotic prescribing among primary care physicians in Switzerland with medium to high antibiotic prescription rates.

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ost antibiotics in human medicine are prescribed in primary care for respiratory tract and urinary tract infections,1-4 contributing to increased populationlevel antibiotic resistance.5-7 Effective strategies to reduce antibiotic prescribing in primary care, like face-to-face education or communication training for primary care physicians, are resource intense, costly, challenging to apply on a large scale, and may not reach clinicians with high prescription rates who are not motivated to participate in such interventions.^{8,9} Therefore, system-wide strategies to improve antibiotic use in primary care are needed. Peer comparison audit and feedback can be an effective tool to modify physician behavior and be applied on a health system level as a low-cost intervention for reducing antibiotic prescribing in primary care.¹⁰⁻¹² Only a few randomized clinical trials have evaluated audit and feedback interventions for antibiotic prescribing in primary care at the health system level. These trials have produced inconsistent results,^{10,12-14} which may be owing to the intensity and type of feedback interventions, different settings, and baseline level of prescribing.

A nationwide pilot trial in Switzerland that investigated the feasibility and effectiveness of quarterly personalized prescription feedback to primary care physicians based on aggregated practice-based claims data did not observe a reduction in antibiotic prescribing in the intervention group.¹¹ Feedback on that trial was based on aggregated information and did not allow ascribing antibiotic prescriptions to individual patients. Thus, patient factors, such as comorbidities, that may influence prescribing behavior and adverse consequences of the intervention, could not be adequately considered.

The main objective of the present randomized clinical trial was to investigate the effect of patient-level claims data audit and feedback with peer benchmarking provided to physicians compared with no intervention on antibiotic prescribing in primary care. Adverse health outcomes (all-cause hospitalizations and infection-related hospitalizations) were a key secondary end point. An additional goal was to explore the feasibility of developing a nationwide antibiotic prescribing monitoring program using claims data.

Methods

Study Design

This pragmatic randomized clinical trial was conducted from January 1, 2018, to December 31, 2019, among primary care physicians in Switzerland with medium to high antibiotic prescription rates and was based on routinely collected individual claims data of 3 large health insurers in Switzerland (Sanitas, CSS, and Helsana) providing coverage for approximately 50% of Swiss residents of all ages. We used pseudonymized physician and anonymized patient identifiers that were created by data managers of the insurance companies to ensure confidentiality. Claims data were formatted by data managers of the health insurers according to a standard protocol that allowed for data import and the identification of the relevant physician population, the generation of the antibiotic prescribing feedback, and the generation of the full claims data

Key Points

Question Does automated quarterly antibiotic prescribing feedback with peer benchmarking over 2 years reduce antibiotic prescribing in the second year of the intervention among primary care physicians who are the top 75% prescribers of antibiotics (ie, with medium to high antibiotic prescription rates)?

Findings In this randomized clinical trial of 3426 primary care physicians in Switzerland, there was a 4% relative increase in antibiotic prescribing during the second year of the intervention (2019) compared with the baseline year (2017). The median annual antibiotic prescribing rate per 100 consultations was 8.2 in the feedback and audit group and 8.4 in the control group in the second year of the intervention.

Meaning Among primary care physicians with medium to high antibiotic prescription rates, antibiotic prescribing audit and feedback did not reduce antibiotic prescribing.

set for the final analysis (eFigure 1 in Supplement 1). The study protocol was approved by all ethics committees in Switzerland (Leitethikkommission Nordwest- und Zentralschweiz, Project-ID 2017-00888) and need for consent of participating physicians and their patients was waived based on article 34b of the Swiss Federal Act on Research involving Human Beings, which defines the rules for the exceptional use of health data without formal consent of participants. Details of the trial protocol have been previously published¹⁵ (trial protocol in Supplement 2). All analyses followed the Consolidated Standards of Reporting Trials (CONSORT) reporting guideline using intention-to-treat principles based on the final data set, which became available in fall 2020.¹⁶⁻¹⁸ We also conducted a per-protocol analysis and 3 post hoc sensitivity analyses by excluding physicians who were identified as outliers for antibiotic-related end points at baseline, first year of the intervention, and second year of the intervention, by excluding practices of 1 to 2 or more than 3 physicians working under the same license number. The trial is registered at ClinicalTrials.gov (NCT03379194).

Participants

We included board-certified primary care physicians and pediatricians with an individual practicing license number (Zentralregisternummer) identified under a unique address who were among the top 75% prescribers of antibiotics (ie, with medium to high prescription rates) with at least 100 patient contacts per year and identified 3426 from 4888 physicians assessed for eligibility in the claims data set of 2016. License numbers of large group practices and hospital-based ambulatory facilities were excluded, but we included practices where more than 1 physician was working under 1 license number (eTable 10 in **Supplement 1**). Because of the delayed administrative processing of claims data, the 2016 data set was taken for identification of physicians and sample size calculation.

Randomization and Masking

Eligible physicians were randomized to the intervention and control groups in a 1:1 ratio using a computer-generated algorithm in R, version 4.0.2 (R Group for Statistical Computing).¹⁹

Physicians in the intervention groups were formally blinded regarding the fact of being included in an intervention trial, and physicians in the control group were not informed that their antibiotic prescription was monitored for the duration of the trial. As all trial-relevant data were collected by automated processes for claims data by health insurances, the outcome assessment may also be considered formally blinded.

Procedures

Continuously updated quarterly antibiotic prescription feedback contained the personal overall prescription rates and antibiotic type per 100 consultations and year, as well as the personal prescription rates for the 3 months of the same period of the preceding year, with each category compared with the prescription rates of peer physicians. In addition, each mailing contained a message for action (eFigure 2 in Supplement 1).¹⁵ The first mailing was sent on December 22, 2017, and was followed by 7 additional quarterly mailings, with the last mailing sent in late September 2019.

Because there is a 6-month delay in health insurance billing records, processing the yearly prescription feedback was based on the data from 9 months before the respective mailing. The first feedback mailing was based on antibiotic prescription rates between April 2016 and March 2017 (eFigure 3 in Supplement 1). Quarterly prescription feedback was prepared by 1 trial statistician (G.M.) with no further involvement in the analysis, and printed, packaged, and mailed by staff not otherwise involved in the trial.¹⁵

With the first postal mailing, all primary care physicians in the intervention group received an accompanying letter explaining the intervention, a response card for physicians wishing to opt out, and evidence-based guidelines in German and French on antibiotic prescribing for respiratory tract and urinary tract infections.^{11,15,20} With the second feedback mailing, information on antibiotic resistance and its regional distribution from the Swiss Centre for Antibiotic Resistance²¹ was provided. By the end of the study, 65 practices had closed and 53 had withdrawn consent to participate further in the study (eFigure 4 in Supplement 1). All information material and guidelines were also made available to physicians in the intervention group on a password-protected trial website.

Outcomes

The primary outcome was the overall antibiotic prescription rate per 100 consultations in the second year of intervention (long-term intervention effect). The secondary outcomes were (1) overall antibiotic use per 100 patient consultations in the first year (short-term intervention effect) (2) and over 2 years while considering 2 repeated measurements, over the first and the second year of intervention; (3) use of broadspectrum antibiotics (quinolones and oral cephalosporins) per 100 patient consultations; (4) all-cause hospitalizations and infection-related hospitalizations¹⁵; (5) antibiotic use in 3 specific patient age groups (\leq 5 years, 6-65 years, and >65 years). The last 3 secondary outcomes were to be evaluated separately over the second and first years of the intervention. The detail of the calculation of prescription rates is provided in the eAppendix in Supplement 1.

Statistical Analysis

Statistical analysis was conducted from September 1, 2021, to January 31, 2022. Details of the sample size calculation are provided elsewhere.¹⁵ We calculated monthly and annual medians for the number of antibiotic prescriptions per 100 consultations, with associated IQRs, per physician license number (the unit of analysis) for the baseline year (2017) and the intervention years (2018-2019). The primary and secondary outcomes were modeled using analysis of covariance,²² with the intervention as a factor of interest, baseline antibiotic prescription rates in 2017 and comorbidities (immunosuppressionlike conditions, metabolic disorders, cardiovascular disease, neurologic disorders, respiratory disease, and other conditions) as covariates, and an interaction term for the intervention with time. We report the coefficient estimates in percentage change in prescriptions per 100 consultations with the respective 95% CIs.

Effects for the overall time of the intervention (24 months) were assessed with a linear mixed model, including the intervention, time (baseline and 24 months), and an interaction term for the intervention with time. From the model including the same covariates as the primary analysis and randomized physicians as random effects, we derived mean percentage changes from baseline for the intervention and control groups. Because of the skewed nature of antibiotic prescription rates, we log-transformed rates and back-transformed model estimates with log-log model formulations.

The number of hospitalizations (overall and infection related) were modeled using Poisson regressions, including the intervention as a factor of interest and all covariates from the primary analysis. Finally, we conducted stratified analyses by age groups as for the primary analysis over the respective periods (eAppendix in Supplement 1). To report some of the baseline characteristics (**Table**), we used patient-level data, but for modeling the data were aggregated on the physician license number. Statistical analyses were performed using R, version 4.0.2 (R Group for Statistical Computing).¹⁹

Results

Of 4888 physicians assessed for eligibility, 3426 were randomized to the intervention (n = 1713) and control groups (n = 1713), with 1591 and 1579 physicians, respectively, being available for analysis for the second year of the intervention (**Figure 1**). At baseline in 2017, physicians in the intervention group served 629 825 patients and physicians in the control group 622 344 patients and prescribed 212 933 antibiotics in the intervention group and 211 825 antibiotics in the control group, for a total of 2 402 119 consultations in the intervention group and 2 388 406 consultations in the control group. A total of 53 physicians opted out of the intervention (eFigure 4 in Supplement 1). Patients' characteristics were well balanced between study groups (Table).

Median annual antibiotic prescription rates per 100 consultations in the year preceding the trial were 8.4 (IQR, 6.3-11.5) in the intervention group and 8.4 (IQR, 6.4-11.6) in the control group. In the intervention group, prescription rates per 100

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	Participants, No. (%)		
Characteristic	Total (N = 1 252 169 [100])	Intervention (n = 629 825 [50.3])	Control (n = 622 344 [49.7])
No. of patients, median (IQR)	337 (236-480)	338 (233-482.2)	336 (238-475)
Patients' age, y			
≤5	187 557 (15.0)	95 468 (7.6)	92 089 (7.4)
6-64	711 234 (56.8)	358 622 (28.6)	352 612 (28.2)
≥65	353 378 (28.2)	175 735 (14.0)	177 643 (14.2)
Female patients	686 633 (54.8)	344 098 (27.5)	3425 35 (55.0)
Comorbidities ^a			
Immunosuppression-like conditions ^b	1 400 627 (100)	699 822 (55.9)	700 805 (56.0)
Metabolic disorders	1 270 058 (100)	635 431 (50.0)	634 627 (50.0)
Cardiovascular disease and neurologic disorders	2 523 931 (100)	1 260 454 (49.9)	1 263 477 (50.1)
Respiratory disease	834 878 (100)	417 552 (50.0)	417 326 (33.3)
Other ^c	2 836 897 (100)	1 419 263 (50.0)	1 417 634 (50.0)
Consultations			
No. (%)	4 790 525 (100)	2 402 119 (50.1)	2 388 406 (49.9)
Mean (SD)	1447.7 (805.9)	1454.1 (817.1)	1441.4 (794.7)
Median (IQR)	1283 (869-1861)	1301 (872.5-1859.3)	1271 (865-1862)

Table. Baseline Characteristics in Baseline Year 2017

^a Multiple comorbidities

per patient are possible.

^b Classification according to pharmacy cost groups.

^c Psychiatric disorders,

gastric acid-related disorders, and iron deficiency anemia.





consultations were 8.3 (IQR, 6.2-11.7) in the first year and 8.2 (IQR, 6.1-11.4) in the second year; in the control group, prescription rates per 100 consultations were 8.5 (IQR, 6.3-12.0) in the first year and 8.4 (IQR, 6.0-11.8) in the second year. Monthly median prescription rates are provided in **Figure 2**. In comparison with the year 2017 prior to the intervention, there was a 4.2% (95% CI, 3.9%-4.6%) increase in the antibiotic prescription rate during the intervention phase for the entire cohort (eTable 1 in Supplement 1). Relative to these increased prescription rates in 2017, there was a small statisti-

Figure 2. Median Antibiotic Prescription Rates per Month



Error bars indicate IQRs.

cally nonsignificant -0.1% (95% CI, -1.2% to 51.0%) reduction in antibiotic prescriptions per 100 consultations in the feedback group compared with the control group in the second year of the intervention (primary end point) (**Figure 3**). During the first year and the entire trial period, antibiotic prescription rates in the intervention group additionally increased during the first year by 0.5% (95% CI, -0.1% to 1.2%) and during the entire trial period by 0.5% (95% CI, -0.2% to 1.3%) when compared with the control group (**Figure 4**). Findings were similar when restricting the analysis to practices with fewer than 3 physicians working under 1 license number (eTable 8 in Supplement 1).

Prescription rates for specific antibiotics also increased during the intervention period when compared with the baseline year (eTable 2 in Supplement 1). Relative to these increased rates, small relative reductions in antibiotic prescriptions were noted in the intervention group compared with the control group during both years of the intervention (eTable 4 and eTable 6 in Supplement 1); these reductions were not statistically significant with the exception of quinolone prescriptions during the second year of the intervention (-0.9%; 95% CI, -1.5% to -0.4%) (Figure 3). No statistically significant differences in antibiotic prescription rates were noted between the feedback and control groups for all prespecified age-related subgroup analyses (eTables 2, 3, and 5 in Supplement 1). Estimates from the per-protocol analysis likewise showed no reductions in antibiotic prescriptions between both groups (eTable 9 in Supplement 1). Also, no differences in infection-related and overall hospitalization rates were found between both groups during all observation periods (eTable 7 in Supplement 1).

Discussion

In this nationwide insurance claims data-based pragmatic randomized clinical trial, primary care physicians and pediatricians in Switzerland with the highest prescription rates were randomized to receive regular prescribing feedback, the proFigure 3. Change in Prescription Rates per 100 Consultations by Physicians in the Intervention vs the Control Group in the Second Year of the Intervention (Intention-to-Treat Analysis) Relative to the Baseline Year 2017

Prescribing for specific populations or specific antibiotics	Between-group difference, % (95% CI)	Favors intervention	Favors control
All patients			
Base model	-0.1 (-1.2 to 1.0)		
Multivariable model	-0.1 (-1.2 to 0.9)		
Age ≤5 y			
Base model	-0.8 (-2.0 to 0.5)		
Multivariable model	-0.6 (-1.9 to 0.6)		
Age 6-65 y			
Base model	<0.01 (-1.0 to 1.0)		:
Multivariable model	0.01 (-1.0 to 1.0)		
Age >65 y			
Base model	0.4 (-0.6 to 1.4)		
Multivariable model	0.5 (-0.5 to 1.5)		
Macrolides			
Base model	0.04 (-0.5 to 0.6)		
Multivariable model	0.03 (-0.5 to 0.6)		
Othe B-lactams			
Base model	-0.7 (-1.9 to 0.5)		
Multivariable model	-0.7 (-2.0 to 0.5)		
Quinolones			
Base model	-0.9 (-1.5 to -0.4)		
Multivariable model	-1.0 (-1.6 to -0.5)		
	. ,	-2 -1 (
		-2 -1 (n difforence %

The base model is treatment variable adjusted to baseline prescription rate and the multivariable model is adjusted to the predefined comorbidities.

vision of community-based antibiotic resistance data, and evidence-based guidelines for respiratory tract and urinary tract infections vs not receiving any information. For the primary end point, the overall antibiotic prescription rate in the second year of the intervention, a very small—from a public health perspective, not relevant—reduction of antibiotic prescription rates was noted. All prespecified subgroup analyses

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Figure 4. Change in Prescription Rates per 100 Consultations by Physicians in the Intervention vs the Control Group in the First Year of the Intervention (Intention-to-Treat Analysis) Relative to the Baseline Year 2017

Prescribing for specific populations or specific antibiotics	Between-group difference, % (95% CI)	Favors Favors intervention control
All patients		
Base model	0.5 (-0.1 to 1.2)	
Multivariable model	0.5 (-0.1 to 1.2)	
Age ≤5 y		
Base model	-0.3 (-1.5 to 0.9)	
Multivariable model	-0.2 (-1.4 to 1.0)	
Age 6-65 y		
Base model	-0.2 (-0.9 to 0.5)	
Multivariable model	-0.2 (-0.9 to 0.5)	
Age >65 y		
Base model	-0.5 (-1.2 to 0.1)	
Multivariable model	-0.5 (-1.2 to 0.2)	
Macrolides		
Base model	-0.1 (-0.6 to 0.4)	
Multivariable model	-0.1 (-0.6 to 0.4)	
Othe B-lactams		
Base model	-0.6 (-1.7 to 0.6)	
Multivariable model	-0.6 (-1.8 to 0.6)	
Quinolones		
Base model	0.04 (-0.4 to 0.5)	
Multivariable model	-0.03 (-0.5 to 0.5)	
		-2 -1 0 1
		Between-group difference, %

The base model is treatment variable adjusted to baseline prescription rate and the multivariable model is adjusted to the predefined comorbidities.

showed also no difference between the 2 groups, with the exception of a -0.9% difference in quinolone prescriptions between the intervention and control groups during the second year of the intervention.

Several smaller trials that were conducted in selected general practices used antibiotic prescription feedback in combination with personalized expert feedback,²³ academic detailing,⁹ a practice accreditation program,²⁴ or in combination with decision support systems^{12,25,26} and found a relative reduction in antibiotic prescriptions of approximately 5%. Some of these trials recruited motivated primary care physicians already engaged in education programs to reduce antibiotic prescriptions or offered financial incentives for trial participation.²⁵ Only a few trials used routine prescription feedback for clinicians in primary care with high antibiotic prescription rates on the health system level addressing the entire primary care physician community. In a trial in the UK, a 3.3% reduction in antibiotic prescriptions over 6 months was found when a single letter from England's chief medical officer with information on high prescription rates was sent to the top 20% of antibiotic prescribers.¹⁰ In a Canadian trial, the same intervention approach was used for the 25% top prescribers in Ontario and a nonstatistically significant 3.4% reduction over 12 months was found.¹⁴ In a trial from New Zealand, a single letter providing peer antibiotic comparison found reduced antibiotic use only among clinicians with high prescription rates but not those with low prescription rates.²⁷ Another trial from Australia provided antibiotic prescription feedback twice over 6 months in unselected rural practices and found no difference in antibiotic prescribing.¹³

The negative findings from this trial are in contrast with results from the UK and Canadian trials, which both included system-wide practices. First, antibiotic prescription rates in the UK trial were about 30% higher than in the present trial.^{10,14} The Canadian trial was based on drug sales data with no denominator information. The present trial included-for reasons of sample size-about 80% of eligible practices and represents a more heterogeneous group of practices in terms of antibiotic prescribing patterns. Switzerland has one of the lowest antibiotic consumption rates among European countries.²⁸⁻³⁰ Thus, it seems difficult to further reduce antibiotic prescribing with our chosen approach in a setting with already low prescription rates, although we know that prescriptions of antibiotics in primary care for upper respiratory tract and urinary tract infections are still too high³¹ and the spread of multidrug resistance remains a problem in particular for gram-negative bacteria expressing extendedspectrum β-lactamases.^{21,32}

Limitations and Strengths

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Our trial has several limitations. Due to the long processing time of claims data by health insurers, prescription feedback was sent to physicians with a delay of 6 months, making it likely less relevant or more difficult to interpret in the actual clinical situation.³³ Swiss claims data do not contain any diagnostic information from primary care; therefore, it was not possible to provide feedback on the appropriateness of antibiotic prescriptions. Legal issues in regard to the privacy of health data are major obstacles to overcome this data deficit in Switzerland. Social scientists emphasize the social normative aspects when aiming at behavior change in antibiotic stewardship interventions and the necessity to link peer comparisons appropriately with the top percentage of prescribers.³³ We chose a mean prescription benchmark rather than practices with the top percentage of prescribers as a reference to acknowledge prescribing variations due to differences in the case mix and the lack of diagnostic data for the reason for antibiotic prescriptions. Other investigators have advocated the use of the top percentage of prescribers as a reference standard.³³ Finally, our trial does not cover the entire Swiss population.

A relative increase in antibiotic prescriptions of 4% was noted in both years of the intervention compared with the baseline in 2017. The Swiss Ministry of Health also reported higher rates of influenza-like illnesses and invasive *Haemophilus influenzae* infections during the intervention period of 2018-2019 when compared with 2017.³⁴ This finding illustrates the importance of evaluating feedback interventions for antibiotic prescribing over sufficiently long periods to accurately account for seasonality effects.

The strengths of this trial are the formal blinding of physicians and data analysis, the conduct of intention-to-treat and per-protocol analyses, the completeness and external validity of our data, and the use of nonaggregated patient-level claims data for certain end points. In our statistical approach– contrary to previous trials^{10,14}–we integrated baseline prescription data, abstained from excluding practices with extreme prescription data, and included covariates to address the patient case mix. Because of the large sample, all our estimates include small 95% CIs. Finally, we were able to explore the theoretical harm from the intervention by the analysis of hospitalization data.

Conclusions

In this randomized clinical trial, quarterly detailed claims data-based prescription feedback to primary care phy-

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Critical revision of the manuscript for important intellectual content: Aghlmandi, Halbeisen, Saccilotto, Signorell, Glinz, Moffa, Widmer, Kronenberg, Bielicki, Bucher.

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Conflict of Interest Disclosures: Dr Aghlmandi reported receiving grants from the Swiss National Science Foundation during the conduct of the study; and grants from the Swiss National Science Foundation outside the submitted work. Dr Saccilotto reported receiving personal fees from the Basel Institute for Clinical Epidemiology and Biostatistics as external consultant during the conduct of the study. Dr Glinz reported being employed by Roche Pharma (Schweiz) AG; data collection for this article was completed before his current employment. Dr Kronenberg reported being the head of the Swiss Federal Office of Public Health (SFOPH) Swiss Antibiotic Resistance Center and being financially supported by the SFOPH during the conduct of the study. Dr Bielicki reported being the project lead for the subproject "Swiss Antimicrobial Stewardship Programme" on behalf of Swissnoso, supported by the Federal Office of Public Health under the umbrella of the National Strategy against Antimicrobial Resistance. Dr Bucher reported receiving grants, support for traveling, consultancy fees, and honoraria from Gilead, Bristol Myers Squibb, ViiV Healthcare, Roche, and Pfizer; and serving as the president of the Association Contre le HIV et Autres Infections Transmissibles, for which he has received support for the Swiss HIV Cohort Study from ViiV Healthcare, Gilead, Bristol Myers Squibb, and MSD, outside the submitted work. No other disclosures were reported.

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sicians did not reduce antibiotic prescriptions during a 2-year intervention in the Swiss primary health care setting with already low antibiotic use when compared with other European countries. Whether health system-wide antibiotic stewardship programs with more individually tailored information on the appropriateness of antibiotic prescriptions, eventually combined with individual physician-targeted incentives, might achieve further reductions in antibiotic use should be evaluated in future trials. Such trials, however, will need much more detailed, routinely collected diagnostic and laboratory patient data.

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- Invited Commentary

Simplicity Matters–Overengineering Feedback Can Be Counterproductive

Jeffrey A. Linder, MD, MPH; Craig R. Fox, PhD

Overengineering is the process of solving a problem in an unnecessarily elaborate or complicated manner. Unfortunately, quality improvement programs that give clinicians feedback are often overengineered and, as a result, are weaker than they otherwise could be.

How does this happen? Content experts invested in improving a specific clinical target-for example, decreasing

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inappropriate antibiotic prescribing-typically solicit input from interested stakeholders to design informa-

tional interventions. Experts and stakeholders run through many scenarios to determine feedback that they believe recipients would want. Pilot testing may engage additional persons particularly interested in the issue and lead to more scenarios and the inclusion of more information.

System leaders, understandably concerned about the impact of negative messages on workforce morale, may edit the feedback to be vaguer, so as not to offend. Adding vagueness increases subtlety and the complexity of feedback. In the end, the overengineered, subtle feedback is often too complicated and too vague to have the desired impact on targeted individuals. Design teams often fall prey to what behavioral scientists refer to as the "curse of knowledge," automatically assuming that naive participants will be able to cut through the complexity and vagueness to interpret the message the same way.

Decades of behavioral science research has found that humans (including physicians) act as "cognitive misers" who default to cognitive processes that minimize computational effort. This gives rise to 2 challenges for behavior change interventions. First, many behaviors are driven by habits formed by repetition of behavior (eg, antibiotic prescribing) in the presence of consistent contextual cues (eg, symptom patterns). When habits are strong, they are more resistant to explicit appeals for change. Second, explicit appeals are less potent when they are less salient or too complex to easily process and

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